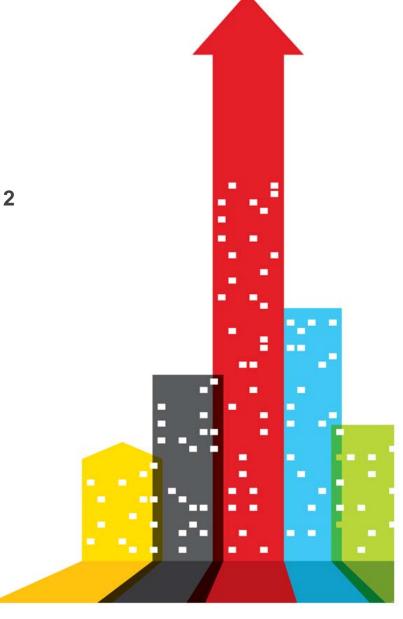


Approximating Correlation Matrices

Mathematical Problems in Industry Workshop 2012

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Roadmap for Correlation Discussion

Themes and Summary of Key Points	
Standard & Poor's Overview	Company profile Quantitative analytics at S&P
Portfolio Modeling and Correlation	Modeling dependence with correlation
Representing Correlation: Factor Models	 Properties of correlation matrices Empirical correlation estimates Block correlation representation Factor models Optimal factor model approximations
Localized Factor Models	 Localized Factor Models Best approximation of localized one factor models Semi-analytic methods for localized one factor models
Open Research Questions	Optimal approximations Computational methods



S&P Overview



S&P Overview

- Standard and Poor's is a leading provider of independent credit analysis, and a key source for financial market intelligence.
- Best known for S&P 500 Index and Credit Ratings
- Wholly owned subsidiary of McGraw-Hill
- About 10,000 employees worldwide with around \$2.9 Billion in combined revenue for Standard & Poor's Ratings Services and S&P Capital IQ.
- Brands include Compustat, CUSIP, Capital IQ, ClariFI, Risk Solutions, GICS, RatingsDirect, IMAKE, R2



Role of Quantitative Analytics Research Group at S&P

Mandate: Support S&P Ratings Services and S&P Capital IQ
with quantitative expertise to grow the businesses and develop
products and services that enhance S&P's position as a
respected voice in the capital markets.

• Priorities:

- Quantitative Support for Ratings Methodology
- Model Development
- Quality and Efficiency
- Quantitative Strategic Vision across Businesses
- Thought Leadership
- Business Advisory and Training



Portfolio Modeling and Correlation



Credit Portfolio Modeling

$$\Pi(t) = \sum_{i=1}^{N} \omega_i \ V_i(t)$$

$$R_{\Pi}\left(T_{0}, T_{H}\right) = \sum_{i=1}^{N} \hat{\omega}_{i} R_{i}\left(T_{0}, T_{H}\right)$$

$$EL = E\left(R_{ref} - R_{\Pi}\right)$$

$$UL = \sigma (R_{ref} - R_{\Pi})$$

$$P(R_{\Pi} < R^*(\alpha)) = \alpha$$



Credit Portfolio Modeling

- Distribution of returns for individual exposures is a key component of determining the portfolio return distribution.
- However, the joint dependence of returns is crucial for determining the portfolio potential for large losses.
- The joint behavior is often modeled by normalizing the distribution of individual returns and estimating a multi-variate distribution for the normalized returns – often referred to as latent variables or asset returns.

$$z_{i} = \Phi^{-1}\left(F_{i}\left(R_{i}\right)\right)$$

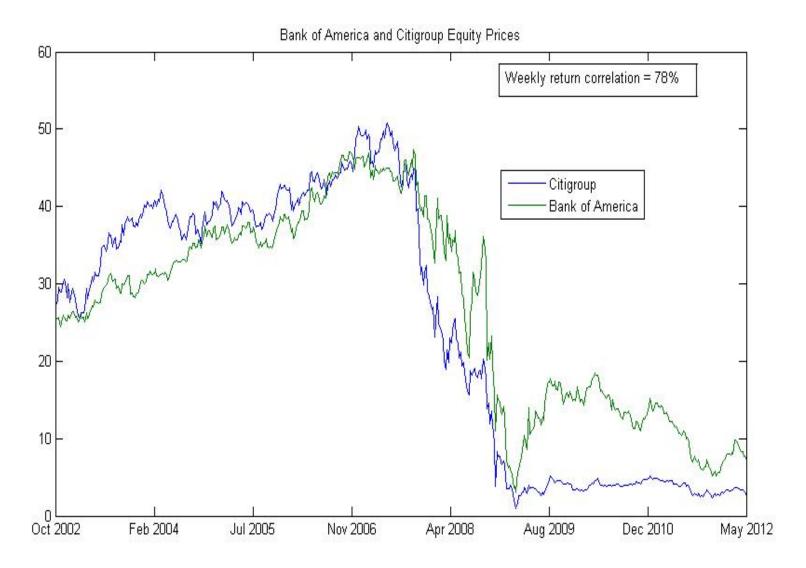
Specifying a correlation matrix (symmetric, pos def with ones on diagonal)

$$P = E\left(zz^{T}\right)$$

or more generally a copula function determines the dependence of the asset returns.



Equity Return Correlation Example





Representing Correlation: Factor Models



Correlation Matrix Properties

- Symmetric
- Positive Definite
- Ones on the diagonal
- Off-diagonal elements represent pair-wise correlation of two assets in the portfolio
- Correlations are usually positive particularly for firm asset returns in credit modeling
- Example: This doesn't work:



Empirical Correlation Estimates

- If time series of 'asset returns' is available or can be proxied by equity returns or other data, pair-wise correlations can be estimated.
- Estimate of full correlation matrix may be somewhat unstable due to missing data and large number of correlation parameters being estimated – N*(N-1)/2 for N firms.
- For large portfolios, describing dependence through a full correlation matrix is impractical.



Factor Models

Gaussian Copula Factor Model – Standard Model

- Normalized 'Asset Return' to Horizon is modeled as a standard Normal random variable.
- Asset return is decomposed into systematic risk component and idiosyncratic risk component.
- Percentage of variance related to systematic risk is 'R-squared'.
- Systematic risk is described by one or more independent standard normal variates, common to all exposures. Each exposure is assigned a set of weights on the factors. The idiosyncratic risk is modeled as standard normal, independent of all factors and other exposures.

$$z_{i} = \sqrt{\rho_{i}} \beta_{i}^{T} \varepsilon_{F} + \sqrt{1 - \rho_{i}} \varepsilon_{I,i} \qquad \beta_{i}^{T} \beta_{i} = 1$$

$$\rho_{ij} = \sqrt{\rho_{i} \rho_{j}} \beta_{i}^{T} \beta_{j}$$

Factor Models

In matrix notation, the factor model is

$$z = \Gamma^{1/2} B \epsilon_F + [I - \Gamma]^{1/2} \epsilon_I$$

The correlation matrix can be represented as

$$P = \Gamma^{1/2} B B^T \Gamma^{1/2} + I - \Gamma$$



Block Correlation Matrix

$$P = \begin{bmatrix} P_{11} & \cdots & P_{1n} \\ \vdots & \ddots & \vdots \\ P_{1n}^T & \cdots & P_{nn} \end{bmatrix} \qquad P_R = \begin{bmatrix} \rho_{11} & \cdots & \rho_{1n} \\ \vdots & \ddots & \vdots \\ \rho_{1n} & \cdots & \rho_{nn} \end{bmatrix}$$

$$P_{ii} = (1 - \rho_{ii}) \mathbf{I}_{N_i} + \rho_{ii} \mathbf{E}_{N_i} \qquad N_i \times N_i$$

$$P_{ij} = \rho_{ij} E_{N_i \times N_j} \qquad N_i \times N_j$$

The reduced correlation matrix P_R is not necessarily positive definite

$$P_{R}^{*} = P_{R} + diag\left(\left[\frac{1-\rho_{11}}{N_{1}}, \dots, \frac{1-\rho_{nn}}{N_{n}}\right]\right) > 0$$



Block Correlation Matrix – Factor Model Representation

- If exposures are classified into groups based on sector, geography, size, etc., and correlations between all members of two groups are assumed the same, the resulting correlation matrix has a block structure.
- The n diagonal blocks are correlation matrices with a constant correlation in each block. Each block has N_i exposures, i = 1 ... n.
- The off-diagonal blocks have identical elements within a block.
- In order to reduce the computation burden of working with large matrices, block correlation matrices can be factored as:

$$z_{ij} = \sqrt{\rho_i + (1 - \rho_i)/N_i} \beta_i^T \varepsilon_F + \sqrt{1 - \rho_i} (\varepsilon_{ij} - \overline{\varepsilon}_i) \quad i = 1 \dots n, \quad j = 1 \dots N_i$$

$$\beta_i^T \beta_i = 1 \quad \beta_i - n \times 1 \text{ vector}$$

 ε_F - $n \times 1$ vector of independent standard Normal variates

 ε_{ij} – iid standard Normal variates

$$\overline{\mathcal{E}}_i = \frac{1}{N_i} \sum_{j=1}^{N_i} \mathcal{E}_{ij}$$



Optimal Factor Model Approximation

- Andersen, Sidenius and Basu (Risk, Nov 2003) show how to find the best K factor approximation to a given correlation matrix P in the sense of minimizing the matrix difference in the Frobenius norm through principal component analysis (PCA).
- Given an estimate of the R-squared diagonal matrix, carry out the following iteration:

$$\begin{aligned} Q_{i}D_{i}Q_{i}^{T} &= P - I + \Gamma_{i} \\ \Gamma_{i+1} &= diag\left(Q_{i}D_{i}^{(K)}Q_{i}^{T}\right) \\ \Gamma_{i} &\to \Gamma_{opt}^{(K)} \\ B_{opt}^{(K)} &= \left(\Gamma_{opt}^{(K)}\right)^{-1/2} Q_{opt}^{(K)} \sqrt{D_{opt}^{(K)}} \end{aligned}$$



Localized Factor Models



Localized One Factor Model

- Assume that there are one global factor and K sector factors.
- A 'localized one factor model' is a factor model where each exposure weights on the global factor and exactly one sector factor.

$$z_{i} = \sqrt{\rho_{i}} \left(\beta_{i}^{0} \varepsilon_{0} + \beta_{i} \varepsilon_{k(i)}\right) + \sqrt{1 - \rho_{i}} \phi_{i}$$
$$\left(\beta_{i}^{0}\right)^{2} + \left(\beta_{i}^{0}\right)^{2} = 1$$

 If all exposures in one sector have the same R-squared value and same factor loadings, then this corresponds to a block correlation matrix.
 However, not every block correlation matrix can be expressed as a localized one factor model.



Localized Two Factor Model

- Assume that there are one global factor and K sector factors and J country factors for a total of (K+J+1) factors.
- A 'localized two factor model' is a factor model where each exposure weights on the global factor, exactly one sector factor and exactly one country factor.

$$z_{i} = \sqrt{\rho_{i}} \left(\beta_{i}^{0} \varepsilon_{0} + \beta_{i}^{S} \varepsilon_{k(i)}^{S} + \beta_{i}^{C} \varepsilon_{j(i)}^{C} \right) + \sqrt{1 - \rho_{i}} \phi_{i}$$
$$\left(\beta_{i}^{0} \right)^{2} + \left(\beta_{i}^{S} \right)^{2} + \left(\beta_{i}^{C} \right)^{2} = 1$$

 If all exposures in one sector have the same R-squared value and same factor loadings, then this corresponds to a block correlation matrix.
 However, not every block correlation matrix can be expressed as a localized two factor model.



Optimal Localized One Factor Approximation

- For a block correlation matrix P with K diagonal blocks and homogeneous correlations within each block, let P_k be the diagonal block correlations and P_{ik} be the off-diagonal block correlations.
- For each exposure in sector k (k = 1 ... K), set the R-squared value to ρ_k
- Construct the correlation matrix corresponding to the ho_{jk} . Find the best single factor approximation using PCA technique. Set

$$\beta_k^0 = \sqrt{\frac{\Gamma_{opt,k}^{(1)}}{\rho_k}} \qquad \beta_k = \sqrt{1 - \left(\beta_k^0\right)^2}$$

If reduced correlation matrix is not positive definite, need to modify.



Semi-Analytic Methods for Localized One Factor Model

For a portfolio with exposures in K sectors, the total loss is

$$L = L_1 + \ldots + L_K$$

 Conditional on the global factor, these sector losses are independent, and each sector is described by a single factor model, so the usual semi-analytic methods can be applied on a sector-by-sector basis:

$$P(L) = E_{\varepsilon_0} \left[\sum_{L_1 + \cdots L_K = L} P(L_1, \cdots, L_K \mid \varepsilon_0) \right]$$

$$P(L_1,\dots,L_K\mid \varepsilon_0) = \prod_{k=1}^K P(L_k\mid \varepsilon_0)$$

$$P(L_{k} \mid \varepsilon_{0}) = E_{\varepsilon_{k}} \left[P(L_{k} \mid \varepsilon_{0}, \varepsilon_{k}) \right]$$

• Alternative: Saddle Point Approximation



Open Research Questions



Optimal Approximations

- For a block correlation matrix, how does the k factor PCA optimal approximation compare with the block correlation factor model representation reduced to k factors?
- For a block correlation matrix, how does the k factor PCA optimal approximation compare with the optimal localized one-factor approximation?
- Can a better localized one-factor model approximation be found that incorporates the number of assets per group?
- How can the optimal localized two factor model be determined?
- Can optimal be measured in terms of portfolio risk metrics as opposed to measured in a matrix norm?



Computational Methods

- For a localized one factor approximation, what are the challenges and limitations of implementing a semi-analytic numerical approach? How efficient is this approach relative to a full Monte Carlo simulation?
- To what extent is it possible to adapt saddle point approximation methods to localized one factor models?



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